

# ECO227Y5 Tutorial 21

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## Question 9.83

**Question:** Suppose that  $Y_1, Y_2, \dots, Y_n$  constitute a random sample from a uniform distribution with probability density function

$$f(y | \theta) = \begin{cases} \frac{1}{2\theta + 1}, & 0 \leq y \leq 2\theta + 1, \\ 0, & \text{otherwise.} \end{cases}$$

(a) Obtain the MLE of  $\theta$ . *- strictly decreasing in  $\theta$*

$$L(y_1, \dots, y_n | \theta) = \prod_{i=1}^n \frac{1}{2\theta + 1} = \frac{1}{(2\theta + 1)^n} \mathbf{I}(y_{(n)} \leq 2\theta + 1) \mathbf{I}(y_{(1)} \geq 0)$$

$$\ln(L(y; \theta)) = \ln\left(\frac{1}{(2\theta + 1)^n}\right) = \overset{=0}{\ln(1)} - n \ln(2\theta + 1)$$

$$\Rightarrow \ln(L(y; \theta)) = -n \ln(2\theta + 1)$$

$$\frac{\partial \ln(L(y; \theta))}{\partial \theta} = \frac{-2n}{2\theta + 1} = 0 \quad \text{does not work!!}$$

Try: want  $\hat{\theta}$  as small as possible, but it can't be too small.

The support must cover all the data so  $y_{(n)} \leq 2\theta + 1 \Rightarrow \hat{\theta}_{MLE} = \frac{y_{(n)} - 1}{2}$

## Question 9.89

**Question:** It is known that the probability  $p$  of tossing heads on an unbalanced coin is either  $\frac{1}{4}$  or  $\frac{3}{4}$ . The coin is tossed twice and a value for  $Y$ , the number of heads, is observed. For each possible value of  $Y$ , which of the two values for  $p$  ( $\frac{1}{4}$  or  $\frac{3}{4}$ ) maximizes the probability that  $Y = y$ ? Depending on the value of  $y$  actually observed, what is the MLE of  $p$ ?

$$Y \sim \text{Bin}(2, p)$$

$$L(Y|p) = \binom{2}{p} p^Y (1-p)^{2-Y}, \quad Y \in \{0, 1, 2\}$$

$$y=0: \binom{2}{0} (1-p)^2 = (1-p)^2 = 1-2p+p^2$$

$$y=1: \binom{2}{1} p(1-p) = 2p-2p^2$$

$$y=2: \binom{2}{2} p^2 = p^2$$

$$y=0: L(0|p) = 1-2p+p^2$$

$$p = \frac{1}{4}: L(0|\frac{1}{4}) = 1 - \frac{2}{4} + \frac{1}{16} = \frac{9}{16} \quad \hat{p}_{MLE} = \frac{1}{4}$$

$$p = \frac{3}{4}: L(0|\frac{3}{4}) = 1 - \frac{6}{4} + \frac{9}{16} = \frac{1}{16}$$

$$y=1: L(1|p) = 2p-2p^2$$

$$p = \frac{1}{4}: L(1|\frac{1}{4}) = \frac{2}{4} - \frac{2}{16} = \frac{3}{8} \quad \text{both equally good!}$$

$$p = \frac{3}{4}: L(1|\frac{3}{4}) = \frac{6}{4} - \frac{18}{16} = \frac{3}{8} \quad \hat{p}_{MLE} = \frac{1}{4} \text{ or } \frac{3}{4}$$

$$y=2: L(2|p) = p^2$$

$$p = \frac{1}{4}: L(2|\frac{1}{4}) = \frac{1}{16} \quad \hat{p}_{MLE} = \frac{3}{4}$$

$$p = \frac{3}{4}: L(2|\frac{3}{4}) = \frac{9}{16}$$

## Question 9.97

**Question:** The geometric probability mass function is given by

$$p(y | p) = p(1 - p)^{y-1}, \quad y = 1, 2, 3, \dots$$

A random sample of size  $n$  is taken from a population with a geometric distribution.

- (a) Find the method-of-moments estimator for  $p$ .
- (b) Find the MLE for  $p$ .

$$a.) \quad E(Y) = \frac{1}{p} = \bar{y} \Rightarrow \hat{p} = \frac{1}{\bar{y}}$$

$$b.) \quad L(y_1, \dots, y_n | p) = \prod_{i=1}^n p(1-p)^{y_i-1} \text{ take } \ln$$

$$\Rightarrow \ln L(y; | p) = \sum_{i=1}^n \ln(p) + (y_i - 1) \ln(1-p)$$

$$\frac{\partial \ln L(y; | p)}{\partial p} = \frac{n}{p} - \frac{\sum_{i=1}^n (y_i - 1)}{1-p} = 0 \Rightarrow \frac{n - np - p \sum_{i=1}^n (y_i - 1)}{p - p^2} = 0$$

$$\Rightarrow n - np - p \sum_{i=1}^n (y_i - 1) = 0$$

$$\Rightarrow n = p \left( n + \sum_{i=1}^n (y_i - 1) \right)$$

$$\Rightarrow p = \frac{n}{n + \sum_{i=1}^n (y_i - 1)} = \frac{n}{n + \sum_{i=1}^n y_i - n} = \frac{n}{\sum_{i=1}^n y_i} = \frac{1}{\bar{y}}$$



# Question 10.6

**Question:** We are interested in testing whether or not a coin is balanced based on the number of heads  $Y$  on 36 tosses of the coin ( $H_0 : p = 0.5$  versus  $H_a : p \neq 0.5$ ). If we use the rejection region  $|y - 18| \geq 4$ , what is

(a) the value of  $\alpha$ ?

(b) the value of  $\beta$  if  $p = 0.7$ ?

Type I error  
 $\alpha = P(\text{Reject } H_0 | H_0 \text{ true})$

a.)  $|Y - 18| \geq 4 \Rightarrow Y - 18 \geq 4$  or  $Y - 18 \leq -4$       *can use z.*  
 $\Rightarrow Y \geq 22$  or  $Y \leq 14$        $Y \sim \text{Bin}(36, p)$   
 $P(Y \geq 22 \text{ or } Y \leq 14 | P=0.5) = \alpha$        $\mu = np$   
 $\sigma^2 = np(1-p)$

$$P(Y \geq 22 | p=0.5) + P(Y \leq 14 | p=0.5) = \alpha$$

$$= P\left(Z \geq \frac{22-18}{\sqrt{36 \cdot 0.5 \cdot 0.5}}\right) + P\left(Z \leq \frac{14-18}{\sqrt{36 \cdot 0.5 \cdot 0.5}}\right) = P\left(Z \geq \frac{4}{3}\right) + P\left(Z \leq -\frac{4}{3}\right) = 2P\left(Z \geq \frac{4}{3}\right) = \alpha \Rightarrow \alpha \approx 0.1876$$

b.)  $\beta = P(\text{Fail to reject } H_0 | H_0 \text{ false})$       *Type II error*  
 $\Rightarrow P(15 \leq Y \leq 21 | p=0.7) = \beta$       *fail to reject so*  
 $\Rightarrow -4 \leq Y - 18 < 4 \Rightarrow 14 \leq Y \leq 22$        $|Y - 18| < 4$

$$\Rightarrow P\left(Z \geq \frac{15 - 0.7 \cdot 36}{\sqrt{36 \cdot 0.7 \cdot 0.3}}\right) + P\left(Z \leq \frac{21 - 0.7 \cdot 36}{\sqrt{36 \cdot 0.7 \cdot 0.3}}\right) = \beta$$

$$\Rightarrow P(Z \geq -3.71) + P(Z \leq -1.53) = \beta \Rightarrow \beta = 0.0633$$

## Question 10.7

**Question:** *True or False.* Refer to Exercise 10.6.

- (a) The level of the test computed in Exercise 10.6(a) is the probability that  $H_0$  is true.
- (b) The value of  $\beta$  computed in Exercise 10.6(b) is the probability that  $H_a$  is true.
- (c) In Exercise 10.6(b),  $\beta$  was computed assuming that the null hypothesis was false.
- (d) If  $\beta$  was computed when  $p = 0.55$ , the value would be larger than the value of  $\beta$  obtained in Exercise 10.6(b).

a.) False  $\alpha = P(\text{Reject } H_0 | H_0 \text{ is true})$

b.) False  $\beta = P(\text{Fail to reject } H_0 | H_0 \text{ is false})$

c.) True.

d.) True as it is closer to 0.5

## Question 10.7

**Question:** *True or False.* Refer to Exercise 10.6.

- (e) The probability that the test mistakenly rejects  $H_0$  is  $\beta$ .
- (f) Suppose that RR was changed to  $|y - 18| \geq 2$ .  
- bigger RR
- (i) This RR would lead to rejecting the null hypothesis more often than the RR used in Exercise 10.6.
  - (ii) If  $\alpha$  was computed using this new RR, the value would be larger than the value obtained in Exercise 10.6(a).
  - (iii) If  $\beta$  was computed when  $p = 0.7$  and using this new RR, the value would be larger than the value obtained in Exercise 10.6(b).

e.) False, it is  $\alpha$ .

f.) i.) True, it is bigger now

ii.) True

iii.) False, it decreases as  $\beta$  is related to not rejecting.



## Question 10.17

**Question:** A survey published in the *American Journal of Sports Medicine* reported the number of meters (m) per week swum by two groups of swimmers — those who competed exclusively in breaststroke and those who competed in the individual medley (which includes breaststroke). The summary statistics are given below.

	$\mu_1$	$\mu_2$
<b>Specialty</b>	<b>Exclusively Breaststroke</b>	<b>Individual Medley</b>
Sample size	130	80
Sample mean (m)	9017	5853
Sample SD (m)	7162	1961
Population mean	$\mu_1$	$\mu_2$

Is there sufficient evidence that the average meters per week practicing breaststroke is greater for exclusive breaststrokers than for individual medley swimmers?

(a) State the null and alternative hypotheses.

$$H_0: \mu_1 - \mu_2 = 0$$

$$H_1: \mu_1 - \mu_2 > 0$$


# Question 10.17

## Question:

Specialty	Exclusively Breaststroke	Individual Medley
Sample size	<u>130</u>	<u>80</u>
Sample mean (m)	9017	5853
Sample SD (m)	7162	1961
Population mean	$\mu_1$	$\mu_2$

Is there sufficient evidence that the average meters per week practicing breaststroke is greater for exclusive breaststrokers than for individual medley swimmers?

- (b) What is the appropriate rejection region for an  $\alpha = 0.01$  level test?
- (c) Calculate the observed value of the appropriate test statistic.

b.)   $\alpha = 0.01$   $n_1 = 130, n_2 = 80$   
 $P(Z \geq \alpha) = 0.01 \Rightarrow \alpha \approx 2.33$

c.) 
$$Z = \frac{\bar{y}_1 - \bar{y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} = \frac{9017 - 5853}{\sqrt{\frac{7162^2}{130} + \frac{1961^2}{80}}} = 9.76$$

# Question 10.17

## Question:

Specialty	Exclusively Breaststroke	Individual Medley
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Population mean	$\mu_1$	$\mu_2$

Is there sufficient evidence that the average meters per week practicing breaststroke is greater for exclusive breaststrokers than for individual medley swimmers?

(d) What is your conclusion?

(e) What is a practical reason for the conclusion in part (d)?

d.) Since test stat yielded  $z = 4.76$  and it is bigger than  $2.33$  we reject  $H_0$  and accept  $H_2$ .

e.) Part D tells us we are 99% confident that the average number of meters per week spent practicing breaststroke is higher for exclusive breaststroke swimmers. As swimmers who compete in individual medley have to practice other styles.

## Question 10.19

**Question:** The output voltage for an electric circuit is specified to be 130. A sample of 40 independent readings on the voltage for this circuit gave a sample mean of 128.6 and standard deviation of 2.1. Test the hypothesis that the average output voltage is 130 against the alternative that it is less than 130. Use a test with level  $\alpha = 0.05$ .

$$\bar{y} = 128.6, n = 40$$

$$sd = 2.1$$

$$H_0: \mu = 130$$

$$H_2: \mu < 130$$

$$\text{test stat: } z = \frac{128.6 - 130}{\sqrt{\frac{2.1^2}{40}}} = -4.22$$

$$P(z \leq \alpha) = 0.05 \Rightarrow \alpha = -1.645$$

since  $z = -4.22 < -1.645$  so reject  $H_0$

and accept  $H_2$ .